

Math 154: Probability Theory, HW 9

DUE APRIL 16, 2024 BY 9AM

Remember, if you are stuck, take a look at the lemmas/theorems/examples from class, and see if anything looks familiar.

1. GETTING OUR HANDS ON BROWNIAN MOTION

1.1. **A computation.** Consider the integral $\int_0^t \mathbf{B}_s^2 ds$.

- (1) Compute $\mathbb{E} \int_0^t \mathbf{B}_s^2 ds$.
- (2) Compute $\mathbb{E} \left| \int_0^t \mathbf{B}_s^2 ds \right|^2$. (*Hint:* as in class, square the integral to get a double integral over $0 \leq r \leq s \leq t$. For $r \leq s$, it may then help to write $\mathbf{B}_s^2 \mathbf{B}_r^2 = (\mathbf{B}_s - \mathbf{B}_r + \mathbf{B}_r)^2 \mathbf{B}_r^2 = (\mathbf{B}_s - \mathbf{B}_r)^2 \mathbf{B}_r^2 + 2(\mathbf{B}_s - \mathbf{B}_r) \mathbf{B}_r^3 + \mathbf{B}_r^4$. Now use independence of increments and knowledge of the distribution of increments.)
- (3) Deduce the variance of $\int_0^t \mathbf{B}_s^2 ds$.

Solution. (1) We have $\mathbb{E} \int_0^t \mathbf{B}_s^2 ds = \int_0^t \mathbb{E} \mathbf{B}_s^2 ds = \int_0^t s ds = \frac{1}{2} t^2$.

(2) We have

$$\begin{aligned} \mathbb{E} \left| \int_0^t \mathbf{B}_s^2 ds \right|^2 &= \int_0^t \int_0^t \mathbb{E} [\mathbf{B}_s^2 \mathbf{B}_r^2] dr ds = 2 \int_0^t \int_0^s \mathbb{E} [\mathbf{B}_s^2 \mathbf{B}_r^2] dr ds \\ &= 2 \int_0^t \int_0^s \mathbb{E} [(\mathbf{B}_s - \mathbf{B}_r)^2 \mathbf{B}_r^2] dr ds + 4 \int_0^t \int_0^s \mathbb{E} [(\mathbf{B}_s - \mathbf{B}_r) \mathbf{B}_r^3] dr ds \\ &\quad + 2 \int_0^t \int_0^s \mathbb{E} [\mathbf{B}_r^4] dr ds. \end{aligned}$$

The second term in the last expression is zero by independence and mean-zero of increments. Since $\mathbf{B}_r \sim N(0, r)$ and $\mathbf{B}_s - \mathbf{B}_r \sim N(0, s - r)$, by independent of increments, we have

$$\begin{aligned} 2 \int_0^t \int_0^s \mathbb{E} [(\mathbf{B}_s - \mathbf{B}_r)^2 \mathbf{B}_r^2] dr ds &= 2 \int_0^t \int_0^s (s - r) r dr ds \\ &= 2 \int_0^t \int_0^s s r dr ds - 2 \int_0^t \int_0^s r^2 dr ds \\ &= \int_0^t s^3 ds - \frac{2}{3} \int_0^t s^3 ds = \frac{1}{4} t^4 - \frac{1}{6} t^4 = \frac{1}{12} t^4, \\ 2 \int_0^t \int_0^s \mathbb{E} [\mathbf{B}_r^4] dr ds &= 6 \int_0^t \int_0^s r^2 dr ds = 2 \int_0^t s^3 ds = \frac{1}{2} t^4. \end{aligned}$$

By combining the previous two displays, we get $\mathbb{E} \left| \int_0^t \mathbf{B}_s^2 ds \right|^2 = \frac{7}{12} t^4$.

- (3) By parts (1) and (2), we have $\text{Var} \int_0^t \mathbf{B}_s^2 ds = \frac{7}{12} t^4 - \frac{1}{4} t^4 = \frac{1}{3} t^4$.

□

1.2. **Brownian Gambler's ruin** (*Hint: use optional stopping!*) Let \mathbf{B} be Brownian motion, and fix $a, b > 0$. Let $\tau_{a,b}$ be the first time τ such that $\mathbf{B}_\tau \in \{-a, b\}$.

(1) Find the probability that $\mathbf{B}_{\tau_{a,b}} = a$.

(2) Compute $\mathbb{E}\tau_{a,b}$.

Solution. (1) By optional stopping and the martingale property of \mathbf{B} , we have $\mathbb{E}\mathbf{B}_{\tau_{a,b}} = 0$.

But $\mathbb{E}\mathbf{B}_{\tau_{a,b}} = -a\mathbb{P}[\mathbf{B}_{\tau_{a,b}} = -a] + b\mathbb{P}[\mathbf{B}_{\tau_{a,b}} = b] = -a\mathbb{P}[\mathbf{B}_{\tau_{a,b}} = -a] + b(1 - \mathbb{P}[\mathbf{B}_{\tau_{a,b}} = -a])$. Thus, we get $\mathbb{P}[\mathbf{B}_{\tau_{a,b}} = -a] = \frac{b}{a+b}$.

(2) By optional stopping and the martingale property of $\mathbf{B}_t^2 - t$, we have $\mathbb{E}\mathbf{B}_{\tau_{a,b}}^2 = \mathbb{E}\tau_{a,b}$.

By part (1), we have $\mathbb{E}\tau_{a,b} = \mathbb{E}\mathbf{B}_{\tau_{a,b}}^2 = a^2\frac{b}{a+b} + b^2\frac{a}{a+b} = \frac{a^2b+ab^2}{a+b}$.

□

1.3. Moment generating function of Gaussians, Brownian motion style. Consider the process $\mathbf{M}_t := \exp\{\lambda \mathbf{B}_t - \mu t\}$, where $\lambda, \mu \in \mathbb{R}$.

- (1) Fix $\lambda \in \mathbb{R}$. For which $\mu = \mu(\lambda) \in \mathbb{R}$ does \mathbf{M} satisfy the martingale property? ($\mu(\lambda)$ will depend on λ .) In what follows, we will always take \mathbf{M}_t for this choice of $\mu = \mu(\lambda)$.
- (2) Fix $\lambda \in \mathbb{R}$. Show that $\mathbb{E}\mathbf{M}_1 = 1$.
- (3) Deduce that if $Z \sim N(0, 1)$, then $\mathbb{E}e^{\lambda Z} = e^{\lambda^2/2}$. (*Hint*: recall $\mathbf{B}_1 \sim N(0, 1)$.)

Solution. (1) As shown in class, for \mathbf{M}_t to be a martingale, we need to find μ such that

$$\left(\partial_t + \frac{1}{2}\partial_x^2\right) \exp\{\lambda x - \mu t\} = 0.$$

The LHS is equal to $\exp\{\lambda x - \mu t\}(-\mu + \frac{1}{2}\lambda^2)$. Thus, it suffices to take $\mu = \frac{1}{2}\lambda^2$.

- (2) By the martingale property, we have $\mathbb{E}\mathbf{M}_1 = \mathbb{E}\mathbf{M}_0 = 1$.
- (3) By part (2), we have $\mathbb{E}e^{\lambda \mathbf{B}_1 - \lambda^2/2} = 1$. Thus, $\mathbb{E}e^{\lambda \mathbf{B}_1} = e^{\lambda^2/2}$. Conclude by the hint. \square

1.4. Ergodicity of the OU process. Suppose X_t is an OU process with initial condition X_0 , that is $dX_t = -X_t dt + dB_t$, where B_t is a Brownian motion.

- (1) Show that $N(0, 1)$ is an invariant distribution for the OU process (see the notes for what this means).
- (2) Let Z_t be an OU process with initial condition $Z_0 \sim N(0, 1)$. That is, $dZ_t = -Z_t + dB_t$, where B is the *same* Brownian motion from above. Define $Y_t = X_t - Z_t$. Show that $Y_t = Y_0 e^{-t}$ for all $t \geq 0$. Deduce that $Y_t \rightarrow 0$ as $t \rightarrow \infty$. (*Hint*: compute the differential equation solved by Y_t using the SDEs for X_t, Z_t ; you can use that any solution to $f'(t) = -f(t)$ is given by $f(t) = f(0)e^{-t}$.)

Solution. (1) As shown in class, it suffices to show that

$$\frac{d^2}{dx^2} p(x) + \frac{d}{dx} (xp(x)) = 0,$$

where $p(x)$ is the pdf for $N(0, 1)$. We check this directly:

$$\begin{aligned} \frac{d}{dx} p(x) &= \frac{d}{dx} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} = -\frac{1}{\sqrt{2\pi}} x e^{-\frac{x^2}{2}}, \\ \frac{d^2}{dx^2} p(x) &= \frac{d}{dx} \left(-\frac{1}{\sqrt{2\pi}} x e^{-\frac{x^2}{2}} \right) = -\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} + \frac{1}{\sqrt{2\pi}} x^2 e^{-\frac{x^2}{2}}. \end{aligned}$$

Thus,

$$\begin{aligned} \frac{d^2}{dx^2} p(x) + \frac{d}{dx} (xp(x)) &= -\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} + \frac{1}{\sqrt{2\pi}} x^2 e^{-\frac{x^2}{2}} - \frac{1}{\sqrt{2\pi}} x^2 e^{-\frac{x^2}{2}} + \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \\ &= 0, \end{aligned}$$

so we are done.

- (2) We have $dY_t = dX_t - dZ_t = -X_t dt + Z_t dt = -Y_t dt$. Now use the hint to get $Y_t = Y_0 e^{-t}$.

□

1.5. Brownian bridge. The Brownian bridge is a “Brownian motion conditioned to hit 0 at time 1”. The point of this exercise is to make this precise in a more natural way.

Let $\{z_k\}_{k=1}^{\infty}$ be a collection of i.i.d. $N(0, 1)$ random variables. For any $N > 0$, define

$$\mathbf{Z}_t^{(N)} := \sum_{k=1}^N \frac{z_k \sqrt{2}}{k\pi} \sin(k\pi t).$$

Show that $\mathbf{Z}_0^{(N)} = \mathbf{Z}_1^{(N)} = 0$. Show that $\mathbb{E}\mathbf{Z}_t^{(N)} = 0$ and that

$$\mathbb{E}|\mathbf{Z}_t^{(N)} - \mathbf{Z}_t^{(M)}|^2 \rightarrow_{N, M \rightarrow \infty} 0.$$

Solution. We know that $\sin(k\pi) = 0$ for any integer k , so $\mathbf{Z}_0^{(N)}, \mathbf{Z}_1^{(N)} = 0$ follows. Since z_k have expectation 0, by linearity of expectation, we have $\mathbb{E}\mathbf{Z}_t^{(N)} = \sum_{k=1}^N \frac{\mathbb{E}[z_k] \sqrt{2}}{k\pi} \sin(k\pi t) = 0$. Moreover, we have

$$\mathbf{Z}_t^{(N)} - \mathbf{Z}_t^{(M)} = \sum_{k=N+1}^M \frac{z_k \sqrt{2}}{k\pi} \sin(k\pi t).$$

Since z_k are i.i.d. $N(0, 1)$, we have

$$\mathbb{E}|\mathbf{Z}_t^{(N)} - \mathbf{Z}_t^{(M)}|^2 = \sum_{k=N+1}^M \frac{2\mathbb{E}|z_k|^2}{k^2\pi^2} \sin^2(k\pi t) \leq \sum_{k=N+1}^M \frac{2}{k^2\pi^2},$$

which is $\leq CN^{-1}$ for some constant $C > 0$. Since $N^{-1} \rightarrow 0$ as $N \rightarrow \infty$, we are done. \square